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1    Temporal variability of ground-shaking and stress  
2    drop in central Italy: a hint for fault healing?

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## 18 Abstract

19 Ground Motion Prediction Equations (GMPEs) are calibrated to predict the  
20 intensity of ground shaking at any given location, based on earthquake mag-  
21 nitude, source-to-site distance, local soil amplifications and other parame-  
22 ters. GMPEs are generally assumed to be independent of time; however,  
23 evidence is increasing that large earthquakes modify the shallow soil condi-  
24 tions and those of the fault zone for months or years. These changes may  
25 affect the intensity of shaking and result in time-dependent effects that can  
26 potentially be resolved analyzing between-event residuals (residuals between  
27 observed and predicted ground motion for individual earthquakes averaged  
28 over all stations). Here we analyze a data set of about 65000 recordings for  
29 about 1400 earthquakes in the moment magnitude range 2.5-6.5 occurred in  
30 central Italy from 2008 to 2017 to capture the temporal variability of the  
31 ground shaking at high frequency. We first compute for each earthquake  
32 between-event residuals in the Fourier domain with respect to a ground mo-  
33 tion prediction equation developed ad-hoc for the analyzed data set. The  
34 between-events show large changes after the occurrence of mainshocks such  
35 as the Mw 6.3, 2009 L' Aquila, the 2016, Mw 6.2 Amatrice and Mw 6.5

36 Norcia earthquakes. Within the time span of a few months after the main-  
 37 shocks, the between-event contribution to the ground shaking varies by a  
 38 factor 7. In particular, we find a large drop in the between-events in the  
 39 aftermath of the l'Aquila earthquake, followed by a slow positive trend that  
 40 leads to a recovery interrupted by a new drop at the beginning of 2014.  
 41 We also quantify the frequency-dependent correlation between the Brune  
 42 stress-drop  $\Delta\sigma$  and the between-events. We find that the temporal changes  
 43 of  $\Delta\sigma$  resemble those of the between-event residuals; in particular, during  
 44 the period when the between-events show the positive trend, the average  
 45 logarithm of  $\Delta\sigma$  increases with an annual rate of 0.19 (i.e., the amplifica-  
 46 tion factor for  $\Delta\sigma$  is 1.56 per year). Breakpoint analysis located a change in  
 47 the linear trend coefficients of  $\Delta\sigma$  versus time in February 2014 although no  
 48 large earthquakes occurred at that time. Finally, the temporal variability  
 49 of  $\Delta\sigma$  mirrors the relative seismic velocity variations observed in previous  
 50 studies for the same area and period, suggesting that both crack-healing  
 51 along the main fault system and healing of micro-cracks distributed at shal-  
 52 low depths throughout the surrounding region might be necessary to explain  
 53 the wider observations of post-earthquake recovery.

## 54 Introduction

55 The intensity of seismic shaking at a given site is a function of the earthquake  
 56 size, style of faulting, source-to-site distance and site condition. Ground  
 57 motion prediction equations (GMPEs) incorporate these functional depen-  
 58 dencies and are time-independent, meaning that the intensity of shaking is

59 assumed to be independent of any process affecting the fault zone. How-  
 60 ever, some observations suggest that shaking intensity changes depending  
 61 on the timing of earthquakes within sequences. For example, it has been  
 62 observed that aftershocks generate lower median ground motion in the high  
 63 frequencies (e.g., Boore and Atkinson 1989) and, therefore, they have been  
 64 flagged in the recent NGA-West2 strong motion data set (Wooddell and  
 65 Abrahamson, 2014) allowing GMPE developers to treat them differently  
 66 from mainshocks. Some of the variability in the observed shaking may be  
 67 explained by variability in the stress drop (e.g., Wu and Chapman, 2017).  
 68 Analyzing the NGA-West2 data, Baltay and Hanks (2014) found that av-  
 69 erage stress drop for mainshocks is 30% larger than the average stress drop  
 70 for aftershocks.

71 Identifying repeating deviations from the medial model can help distinguish  
 72 processes that are not modeled but contribute significantly to ground mo-  
 73 tion variability. For example, several studies correlated earthquake-specific  
 74 residuals (also called between-event or inter-event) to stress drop variability  
 75 (e.g., Anderson and Lei, 1994; Bindi et al., 2007; Cotton et al., 2013; Oth et  
 76 al., 2017; Bindi et al., 2017; Baltay et al., 2017; Ameri et al., 2017; Trugman  
 77 and Shearer, 2018). Therefore, presence of any temporal pattern in the dis-  
 78 tribution of residuals can be used as diagnostic of time-dependent fault or  
 79 medium properties, and ultimately help understanding how non-stationary  
 80 processes, such as protracted seismic sequences or the long precursory phase  
 81 of large earthquakes, affect ground motion. For example, Socquet et al  
 82 (2016) and Piña-Valdés et al. (2018), in discussing the ground-shaking  
 83 time-dependencies observed during the 2014 Iquique subduction sequence,

84 suggested that the temporal changes in the between-event residuals were  
 85 associated with aseismic slip around the rupture area.  
 86 Here we take advantage of a large data set from central Italy to investigate  
 87 the temporal changes of the between-event residuals and their link with the  
 88 stress drop  $\Delta\sigma$  variability. The data set includes about 1400 earthquakes  
 89 in the magnitude range 2.5-6.5, belonging to the main sequences of the last  
 90 ten years, i.e. the 2009 L'Aquila (Chiaraluce et al., 2011) and the 2016-17  
 91 Amatrice-Visso-Norcia (Chiaraluce et al., 2017) sequences. About 60 earth-  
 92 quakes occurred in area of the 2013-14 Gubbio swarm (De Gori et al., 2015)  
 93 are included as well. We first calibrate an ad-hoc GMPE for the Fourier  
 94 amplitude spectra and we evaluate the between-event residuals at different  
 95 frequencies. Then, the between-event are used as exploratory tool to detect  
 96 event-dependent temporal changes in the ground shaking and we conclude  
 97 presenting the temporal variability of  $\Delta\sigma$ .

## 98 Data

99 In this study, we analyze about 65000 recordings (for each component of  
 100 motion) from 1400 earthquakes recorded by 340 stations installed in Central  
 101 Italy (Figure 1 and Figure S1 of the electronic supplements to this article).  
 102 The earthquakes cover the magnitude range from 2.5 to 6.5 and hypocentral  
 103 distances from 10 to 180 km are considered. The data set includes the main  
 104 sequences occurred in the area in the last 10 years, namely the 2009, Mw  
 105 6.3 L'Aquila (indicated as  $e_1$  in Figure 1); the 2016, Mw 6.1 Amatrice ( $e_2$   
 106 in Figure 1); the 2016 Mw 6.1, Visso ( $e_4$  in Figure 1); the 2016, Mw 6.5,

107 Norcia ( $e_3$  in Figure 1) (for a map with the time evolution of the events, see  
 108 Figure S2 in the electronic supplements to this article). Following Bindi et  
 109 al. (2018), in this study we consider the moment magnitude from the Geofon  
 110 catalog for all events with  $M_w \geq 5.7$  but the 2009, L'Aquila mainshock, for  
 111 which we use the Global Centroid Moment Tensor (GCMT) value since the  
 112 Geofon solution is not available (see the Data and Resources section). The  
 113 data set also includes recordings from 59 earthquakes with magnitude larger  
 114 than 2.5 occurred in the area of the 2013-14 Gubbio swarm; a complete  
 115 description of the swarm is given by De Gori et al. (2015) and Valoroso et  
 116 al. (2017). The station distribution is shown in Figure S1 in the electronic  
 117 supplement to this article.

118 We analyse the Fourier amplitude spectra (FAS) of S-waves windows  
 119 band-pass filtered with a variable high pass corner frequency depending on  
 120 the signal-to-noise ratio. The Butterworth high-pass corner varies in the  
 121 range 0.05–0.4 Hz while the low pass one was fixed to 40 Hz. The FAS are  
 122 smoothed using the Konno and Ohmachi (1998) algorithm (the smoothing  
 123 parameter  $b$  was set to 40). Details about the data selection and processing  
 124 are provided by Pacor et al. (2016) and by Bindi et al. (2017).

## 125 **Source parameters**

126 For each earthquake, we consider the source parameter (i.e., stress drop and  
 127 seismic moment) derived by Bindi et al. (2017) using a generalized inversion  
 128 technique (GIT). In the GIT approach (e.g., Castro et al., 1990; Oth et al.,  
 129 2011), the spectral values of a set of earthquakes recorded by a network of

stations are simultaneously inverted to isolate the contribution of source,  
 propagation and site effects. The GIT approach exploits the redundancy  
 of information (that is, the same earthquake is recorded at several stations  
 located at different distances, and several earthquakes are recorded at the  
 same station) to set-up an over-determined system of equations solved in a  
 least-squares sense. To remove unresolved degrees of freedom which generate  
 trade-offs among different components of the solution, some constraints are  
 applied, such as the choice of a reference distance at which the attenuation  
 is assumed to be one and a reference site condition (i.e., one or more stations  
 whose site amplification is assumed to be known). In this study, we use the  
 results of Bindi et al (2018) who applied a non-parametric GIT inversion  
 where any a-priori seismological models for source and attenuation were  
 adopted during the GIT inversion. To estimate the seismic moment and the  
 corner frequency for each earthquake, the resulting non-parametric source  
 spectra were fit to a Brune (1970) source model which assumes a circular  
 fault with uniform stress drop. In Bindi et al. (2018), the source fit was  
 performed allowing a deviation of the high frequency acceleration spectral  
 level from a constant value as predicted by the Brune model. The high-  
 frequency slope of the source spectrum is referred to as  $k_{source}$ . Given the  
 seismic moment and the corner frequency, the stress drop was computed  
 following Eshelby (1957) and Keilis-Borok (1959).

The source parameters are shown in Figure 2, in terms of scaling between  
 $\Delta\sigma$ , Mw and hypocentral depth. Most of the considered depths including  
 those of the mainshocks are located between 5 and 10 km. The stress drop  
 tends to increase with depth and has a strong magnitude dependence (Pacor



et al., 2016; Bindi et al. 2017). The mainshocks have the largest  $\Delta\sigma$ , around 10 MPa. The overall  $\Delta\sigma$  variability covers almost three orders of magnitude. The procedure followed in this study to estimate the uncertainties on  $\Delta\sigma$  is described in the electronic supplement to this article.

## Ground motion model

In this study, we describe the Fourier amplitude spectra  $FAS(f, R)$  at frequency  $f$  of S-waves recorded at hypocentral distance  $R$  with the following seismological model:

$$FAS(f, R) = S(f) \cdot P(f, R) \cdot Z(f) = K \frac{M_0 f^2}{1 + (\frac{f}{f_c})^2} \cdot \frac{1}{R^n} \exp(-\frac{\pi f R}{Q\beta}) \cdot Z(f) \quad (1)$$

where the acceleration source spectra  $S(f)$  is parametrized considering an  $\omega$ -square model (Aki, 1967) and the spectral attenuation with distance  $P(f, R)$  is controlled by the geometrical spreading exponent  $n$  and the anelastic attenuation, the latter being modeled through the quality factor  $Q(f)$ . In equation 1, the constant  $K$  depends on the density and velocity at the source location, on radiation pattern and free surface amplification effects, whereas  $Z(f)$  accounts for site amplification effects. We only consider far-field source terms and extended-source effects are not accounted for. The asymptotic form of the source spectrum is as follow:

$$S(f) \propto \begin{cases} M_0 f_c^2 & \text{if } f \gg f_c \\ M_0 f^2 & \text{if } f \ll f_c \end{cases} \quad (2)$$

172 The source spectrum depends on two parameters, the seismic moment  
 173  $M_0$  and the corner frequency  $f_c$  connected through the stress drop  $\Delta\sigma$   
 174 (Brune, 1970; Eshelby, 1957) as follows:

$$\Delta\sigma \propto M_0 d^{-3} \propto M_0 f_c^3 \quad (3)$$

175 where  $d$  is the source radius. Considering equation 3, equation 2 can be  
 176 re-written as:

$$S(f) \propto \begin{cases} M_0^{\frac{1}{3}} \Delta\sigma^{\frac{2}{3}} & \text{if } f \gg f_c \\ M_0 f^2 & \text{if } f \ll f_c \end{cases} \quad (4)$$

177 If the average stress drop of the analyzed earthquakes is assumed to  
 178 be constant, the scaling of the source spectrum with the earthquake size is  
 179 controlled only by the seismic moment (Aki, 1967). Under this assumption  
 180 and considering a mixed effect regression (Bates et al., 2015), Equations 1  
 181 and 4 suggest the following parametric model for  $FAS(f, R)$ :

$$\ln(FAS) = a_1 + a_2 Mw + a_3 \ln(R) + a_4 R + \delta B_e + \delta B_s + \epsilon \quad (5)$$

where the moment magnitude  $Mw$  is proportional to  $\text{Log}(M_0)$  (Hanks and Kanamori, 1979). In equation 5, the coefficients  $a_i$  are the (frequency dependent) fixed effects which define the median prediction;  $\delta B_e$ ,  $\delta B_s$  are the

random effects for the earthquake and station grouping levels, respectively;  $\epsilon$  is the residual distribution. In order to allow more complex scaling with magnitude, the functional form considered in this study is the following:

$$\ln(FAS) = e_1 + b_1(Mw - M_{ref}) + b_2(M - M_{ref})^2 + [c_1 + c_2(M - M_{ref})]\ln\left(\frac{R}{R_{ref}}\right) + c_3(R - R_{ref}) + \delta B_e + \delta B_s + \epsilon \quad (6)$$

182 with  $M_{ref} = 3.5$  and  $R_{ref} = 1km$ . In equation (6), the fixed effect  
 183 coefficients describe the scaling with distance ( $c_1$  and  $c_3$  are connected to  
 184 the geometrical spreading attenuation and the quality factor, respectively)  
 185 and with magnitude ( $b_1$  and  $b_2$  are controlling the scaling with the seismic  
 186 moment). Coefficient  $c_2$  introduces a magnitude dependency in the attenua-  
 187 tion with distance, while the off-set  $e_1$  depends (at high frequency) on source  
 188 characteristics such as the average stress drop, among other quantities. The  
 189 between-event  $\delta B_e$  quantifies the systematic deviation of recordings for the  
 190 same event with respect the median prediction. At high frequencies, the  
 191 deviation of the stress drop of any earthquakes from the average of the pop-  
 192 ulation is expected to contribute to the  $\delta B_e$  residuals while differences in  
 193 the average radiation pattern among the earthquakes due to uneven station  
 194 distribution can contribute to  $\delta B_e$  at low frequency (along with other fac-  
 195 tors such as differences in the density and velocity at the source location,  
 196 errors in the magnitude values, etc). The between station  $\delta B_s$  random ef-  
 197 fects, sometimes referred to as  $\delta S2S$ , absorb the frequency dependent site  
 198 amplification indicated with  $Z(f)$  in equation 1.

199 The frequency-dependent coefficients of the model (6) and the standard

200 deviations of  $\delta B_e$ ,  $\delta B_s$ , and  $\epsilon$  are listed in Table S1 of the electronic sup-  
 201 plement to this article. The residuals  $\epsilon$  versus hypocentral distance and  
 202  $\delta B_e$  versus magnitude are exemplified in Figure 3 for two frequencies. Over  
 203 the intervals well constrained by data, the average residuals do not show  
 204 systematic trends with the predictor variables but the variability of  $\delta B_e$  in-  
 205 creases with frequency (see also Table S1). Weak trends at short distances  
 206 (at high frequency) and for large magnitude (at low frequencies), which are  
 207 not impacting on the analysis performed in this study, could be removed by  
 208 introducing distance and magnitude hinges in equation (5).

## 209 **Between-event temporal variability**

210 The temporal trend of the between-event  $\delta B_e$  at 10 Hz is shown in Figure  
 211 4 while zooms over different time windows are presented in Figure S3 of  
 212 the electronic supplement to this article. In addition to the large variability  
 213 in the aftermath of the mainshock occurrence, the most striking feature in  
 214 Figure 4 is the positive trend developing from the end of 2009, a few months  
 215 later than the April 6, 2009 L'Aquila mainshock, to late 2013-early 2014.  
 216 In the period from early 2014 to August 2016, when the Amatrice sequence  
 217 started,  $\delta B_e$  shows a large variability with average value close to zero. When  
 218 observed at low frequencies (Figure 4),  $\delta B_e$  shows a weak trend with time.  
 219 As discussed in the model development section,  $\delta B_e$  is expected to absorb,  
 220 at high frequencies, the effect of the stress drop variability. Figure 5 shows  
 221 that the correlation between  $\delta B_e$  and  $\Delta\sigma$  is significant at 10 Hz, whereas the  
 222 correlation is low at 0.75 Hz. It is worth noting that the mainshocks and the

223 largest aftershocks deviate from the average correlation trend defined by the  
 224 aftershock population. We ascribe this behavior to the fact that while  $\Delta\sigma$   
 225 varies over three order of magnitude for small events, it is almost constant  
 226 for earthquakes above magnitude 5. Since GMPE well describe the average  
 227 the ground shaking generated by the largest magnitudes, their  $\delta B_e$  are dis-  
 228 tributed close to zero (Figure 3). The correlation of  $\delta B_e$  with  $\Delta\sigma$  is also  
 229 highlighted in Figure S4 of the electronic supplement where large positive  
 230 residuals are associated to events with  $\Delta\sigma$  higher than 0.6 MPa (i.e., the  
 231 population average; Bindi et al., 2018) while earthquakes with lower stress  
 232 drop have negative residuals. Figure 5 also shows the dependence of  $\delta B_e$  on  
 233 hypocentral depth. The observed trend is reflecting the  $\Delta\sigma$  dependences on  
 234 depth as shown in Figure 2. The degree of correlation measured in terms  
 235 of Pearson coefficient (Figure 6) confirms that the correlation is strongest  
 236 around 10 Hz. The decrease of correlation toward low frequencies reflects  
 237 the diminishing importance of  $\Delta\sigma$  in determining the spectral amplitudes  
 238 at frequencies lower than the corner one while the reduction above 10 Hz  
 239 suggests that source-related effects other than the stress drop also affect the  
 240 ground motion variability at high frequencies. The high frequencies radia-  
 241 tions depends on many factors: small scale slip heterogeneity/slip roughness  
 242 (Causse et al., 2010), rupture velocity and slip source function (Mai et al.,  
 243 2017) and near source attenuation (Purvance and Anderson, 2003). For  
 244 example, analyzing a smaller data set, Bindi et al. (2017) found a correla-  
 245 tion between  $\delta B_e$  and the slope at high frequency of the acceleration source  
 246 spectrum.

## 247 Stress drop temporal variability

248 The stress-drop variability with time (Figure 7a) resembles the variability  
249 observed for  $\delta B_e$ . If the earthquakes are grouped according to the latitude  
250 of their epicenters as shown in Figure 1, and focusing on the average trend,  
251 we observe that:

- 252 •  $\Delta\sigma$  of earthquakes located in the L'Aquila region (Figure 7b) rapidly  
253 diminishes during the first month after the mainshock on April 6; the  
254 recovery starts after about two months (see also Figure 8a);
- 255 • for events located in the Campotosto segment (Chiaraluce et al., 2011)  
256 (Figure 7c), the recovery of the logarithm of  $\Delta\sigma$  develops over a time  
257 span of 4 years, from 2010 to 2013, at an annual rate of 0.17 (i.e., the  
258 amplification factor for  $\Delta\sigma$  is 1.5 per year). We recall that the Cam-  
259 potosto segment includes the northernmost termination of the 2009  
260 sequence and the southern tip of the 2016 fault system (Chiaraluce et  
261 al., 2017). In particular, the four events with magnitude larger than  
262 5 occurred in January 2017 are characterized by large  $\Delta\sigma$  (see Figure  
263 8), as for the largest aftershocks occurred over this segment during the  
264 2009 sequence.
- 265 • a decrease in  $\Delta\sigma$  is observed at the beginning of 2014, although no  
266 large earthquakes occurred at that time. Figure 9 shows the results of  
267 a breakpoint analysis (Bai, 1994; Zeileis et al., 2002; 2003) performed  
268 to detect changes in the coefficients of the linear regression with time.  
269 The analysis identifies a change-point within the period February 10 -

270 March 8, 2014 across which the slope of the logarithm of  $\Delta\sigma$  with time  
 271 reduce from  $4.7e^{-04}$  to  $9.4e^{-05}$  (the amplification factor per year for  
 272  $\Delta\sigma$  reduces from 1.5 to 1). A detailed description of the breakpoint  
 273 analysis is reported in the electronic supplement to this article. The  
 274 cause driving this drop are not known. No large earthquakes occurred  
 275 in the area around February-March 2014; the only notable event is the  
 276 Gubbio swarm (Gualandi et al., 2017). At this stage, it is difficult to  
 277 assess the plausibility of its involvement in the process we are examin-  
 278 ing here. Possible connections with the seismic and aseismic moment  
 279 released during the 2013-14 Gubbio swarm are worth to be the subject  
 280 of future work.

- 281 • earthquakes located in the northern group (Figure 7d) mainly be-  
 282 long to the 2016-17 Amatrice-Norcia-Visso sequence (Chiaraluce et  
 283 al., 2017); also for these events,  $\Delta\sigma$  is larger for the mainshocks and  
 284 for the aftershocks above magnitude 5.5 (Figure S5 in the electronic  
 285 supplement), and decreases after the mainshock occurrence (Figure  
 286 8c). The events occurred in this area before the 2016-17 sequence  
 287 follow the same trends observed for the Campotosto segment.

## 288 **Discussions and conclusions**

289 The between-event residuals  $\delta B_e$  computed for ten years of data in central  
 290 Italy show significant temporal variability at high frequency (Figure 4). On  
 291 one hand, the time dependency of  $\delta B_e$  imply temporal changes of the ground  
 292 shaking that could have an impact over the short term hazard. In the first

couple of months,  $\delta B_e$  at 10 Hz varies in the range from -1 to 1, roughly  
 (i.e., about a factor 0.7 for spectral amplitudes); after a couple of months  
 from L' Aquila mainshock, a trend develops with  $\delta B_e$  increasing on average  
 from about 0 to 0.8, (i.e., about factor 2 in high frequency spectral content).  
 On the other hand, Figure 7 shows that the high-frequency between-events  
 variability resembles the time variability of the stress drop  $\Delta\sigma$ . Tempo-  
 ral variability of  $\Delta\sigma$  has been observed in previous studies. For example,  
 Abercrombie (2014) analyzed 25 earthquakes in three repeating sequences  
 on the San Andreas at Parkfield, observing a long term gradual increase of  
 $\Delta\sigma$  before the 2004 magnitude 6 earthquake. The values show an immedi-  
 ate decrease after the mainshock occurrence before recovering to previous  
 values. Using a long-term stress drop catalog, Chen and Shearer (2013)  
 found relatively stable long-term average stress drop in Southern California  
 but a slow increase trend after large main shocks within the Landers fault  
 zone was also identified, in agreement with a possible long-term fault zone  
 recovery (Li et al., 1998).  
 Fault healing has been shown to promote the generation of high-frequency  
 earthquakes both in laboratory experiments and on natural faults (e.g.,  
 Marone, 1998; McLaskey et al., 2012, Scuderi et al., 2016). The connection  
 between pore pressure and effective normal stress has been also advocated  
 to explain the time variability of the stress drop. Recently, Yoshida et al.  
 (2017) analyzed a swarm triggered by the 2011 Tohoku earthquake, eval-  
 uating temporal changes in stress drop and b-value. They discussed the  
 temporal variations of stress drop (very similar patterns to those observed  
 in this study) in terms of changes in the frictional strength due to fluid



318 migration. In central Italy, pore pressure diffusion due to fluids migration  
 319 played a role in the preparatory phase of the L'Aquila mainshock (e.g., Di  
 320 Luccio et al 2010). However, pore pressure diffusion generally occurs over  
 321 time scales of weeks to months. Thus, we reckon it is difficult to attribute  
 322 solely to migration of fluids or pore pressure the variations we observe, which  
 323 occur over time scales of several years.

324 Among other techniques, monitoring changes in seismic velocities has been  
 325 shown to be effective in detecting fault healing and reloading processes (e.g.,  
 326 Brenguier et al., 2008; Chen et al., 2010). For example, Peng and Ben-Zion  
 327 (2006) investigated the temporal variations of seismic velocity along the  
 328 north Anatolian fault analyzing repeating earthquake clusters in the after-  
 329 shock zones of the 1999 Izmit and Düzce earthquakes. The authors observed  
 330 a sharp seismic velocity reduction immediately after the Düzce main shock,  
 331 followed by a gradual logarithmic-type recovery. They concluded that the  
 332 temporal changes of material properties occur in the topmost portion of  
 333 the crust and, although the change is more prominent at stations located  
 334 close to recently ruptured fault zones, it is not limited to the immediate  
 335 vicinity of the fault zone. In central Italy, Soldati et al (2015) computed  
 336 the relative velocity variation from cross-correlations of noise data over the  
 337 period 2008-2012, including the 2009 L'Aquila main shock. The temporal  
 338 variation obtained for the relative velocity (reproduced in Figure 10) has a  
 339 trend very similar to the stress drop: an abrupt co-seismic decrease at the  
 340 time of the main shock occurrence, followed by an unstable behavior for a  
 341 few months and, finally, a recovery of the velocity (see also in Figure 7c).  
 342 Regarding the spatial distribution of the co-seismic velocity drop, Soldati

et al. (2015) compared the velocities changes averaged over a one-month time window selected before and after the main shock occurrence, and excluding the day of the main shock. They found (see their Figure 5) that the drop was maximum over the area surrounding the L'Aquila epicentre and in the northeast direction from the fault zone, including the Campotosto area. The similarities of the trends observed for the stress drop and for the relative velocity variations suggest that, in agreement with Heckels et al (2018), the recovery can be associated both to crack-healing along the main fault system and to healing of micro-cracks distributed at shallow depths throughout the surrounding region.

## Data and Resources

The R software (R Development Core Team, 2008; <http://www.R-project.org>) has been used in this study to perform the regressions. In particular, the packages lme4 (Bates et al., 2015; <https://cran.r-project.org/web/packages/lme4/news.html>); ggplot (Wickham, 2009; <http://ggplot2.org>); changepoint (Killick, R., & Eckley, I.A., 2014; <https://www.jstatsoft.org/article/view/v058i03>); strucchange (Zeileis et al., 2002; <http://www.jstatsoft.org/v07/i02/>). The waveforms used in this study have been downloaded from European Integrated Data Archive-EIDA (<https://www.orfeus-eu.org/data/eida/>) and from the Italian Civil Protection (DPC) repository (<http://ran.protezionecivile.it/IT/index.php>). Regarding the permanent networks, we used data from the networks with FDSN code: MN, IV, IT (<http://www.fdsn.org/networks/>). The moment magnitude used in this study for all earthquakes larger than 5.7

366 have been taken from the Geofon moment tensor catalog ([http://geofon.gfz-](http://geofon.gfz-potsdam.de/eqinfo/list.php?mode=mt)  
367 [potsdam.de/eqinfo/](http://geofon.gfz-potsdam.de/eqinfo/list.php?mode=mt) [list.php?mode=mt](http://geofon.gfz-potsdam.de/eqinfo/list.php?mode=mt)). Only for L'Aquila mainshock, we  
368 used the GCMT solution ([http://www.globalcmt.org/](http://www.globalcmt.org/CMTsearch.html) [CMTsearch.html](http://www.globalcmt.org/CMTsearch.html)).  
369 The earthquake locations are taken from the INGV bulletin ([http://cnt.rm.ingv.it/](http://cnt.rm.ingv.it/inside)  
370 [inside](http://cnt.rm.ingv.it/inside)). All the Internet sites have last accessed on December 2017. Some of  
371 the Figures were prepared with GMT (Wessel and Smith, 1991).

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378 in Figure 10.

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## Figure captions

**Figure 1.** Map of earthquake epicenters (circles) analyzed in this study (see Data and Resources section). Circles are filled according to the latitude of the epicenter, assuming arbitrary thresholds at latitudes 42.4 and 42.68. A few earthquakes belonging to the 2014 Gubbio swarm are also included. The focal mechanisms of earthquakes with magnitude larger than 6 are shown as beach ball taken from Geofon and from the Global Centroid Moment Tensor catalogs (see Data and Resources section). The rectangles depict the surface projection of the faults as given in Luzi et al. (2016).

**Figure 2.** Scaling relationships between stress drop  $\Delta\sigma$ , hypocentral depth and moment magnitude Mw for the earthquakes analyzed in this study (Bindi et al., 2018). The trend-lines are estimated through a local regression (Loess) performed using the ggplot2 package in R (Wickham, 2009).

**Figure 3.** Observation minus prediction residuals versus predictor variables for the model in equation (1). The residual  $\epsilon$  versus distance and the between event  $\delta B_e$  versus moment magnitude Mw are shown for the regressions performed at 0.75 Hz (left) and 10 Hz (right).

**Figure 4.** (a) Between-event residuals ( $\delta B_e$ ) versus time, at 10 Hz. Earthquakes belonging to the Gubbio swarm (triangles; see Figure 1) are not considered for the evaluating the local-trend analysis; zooms over different time windows are presented in Figure S3 of the electronic supplements to this article. (b) Between-event versus time, at 0.75 Hz. (c) the same as in

580 (b) but zooming over the 2016 sequence. Vertical bars represent the 95%  
581 confidence interval for  $\delta B_e$ .

582 **Figure 5.** Between-event  $\delta B_e$  residuals versus stress drop  $\Delta\sigma$  (top) and  
583 hypocentral depth (bottom), considering the results for 0.75 Hz (left) and  
584 10 Hz (right).

585 **Figure 6.** Person correlation coefficients between  $\Delta\sigma$  and  $\delta B_e$  at different  
586 frequencies.

587 **Figure 7.** Temporal variability of stress drop  $\Delta\sigma$ . (a) complete dis-  
588 tribution of earthquakes; (b) only earthquakes located in proximity of the  
589 2009 L' Aquila mainshock ; (c) only earthquakes occurred in the Campotosto  
590 segment; (d) earthquakes occurred in the area corresponding to the 2016-17  
591 mainshocks. For the location of the earthquakes in panels (b) through (d),  
592 see Figure 1. Vertical bars represent the 95% confidence interval for  $\Delta\sigma$ .  
593 Zooms over different windows are available in Figure 8.

594 **Figure 8.** Temporal variability of stress drop  $\Delta\sigma$ , different zooms of  
595 Figure 7.

596 **Figure 9.** Results of the breakpoint analysis (Zeileis et al., 2002; 2003).  
597 A change in the linear trend with of  $\Delta\sigma$  (dashed lines) is detected between  
598 February 10 and March 8, 2014 (vertical gray line). Details of the analysis  
599 are reported in the electronic supplements to this article.

600 **Figure 10.** Comparison between the relative shear wave velocity vari-  
601 ation computed by Soldati et al (2015) (points) and the  $\Delta\sigma$  (squares) time  
602 variability.

603 **Figures**

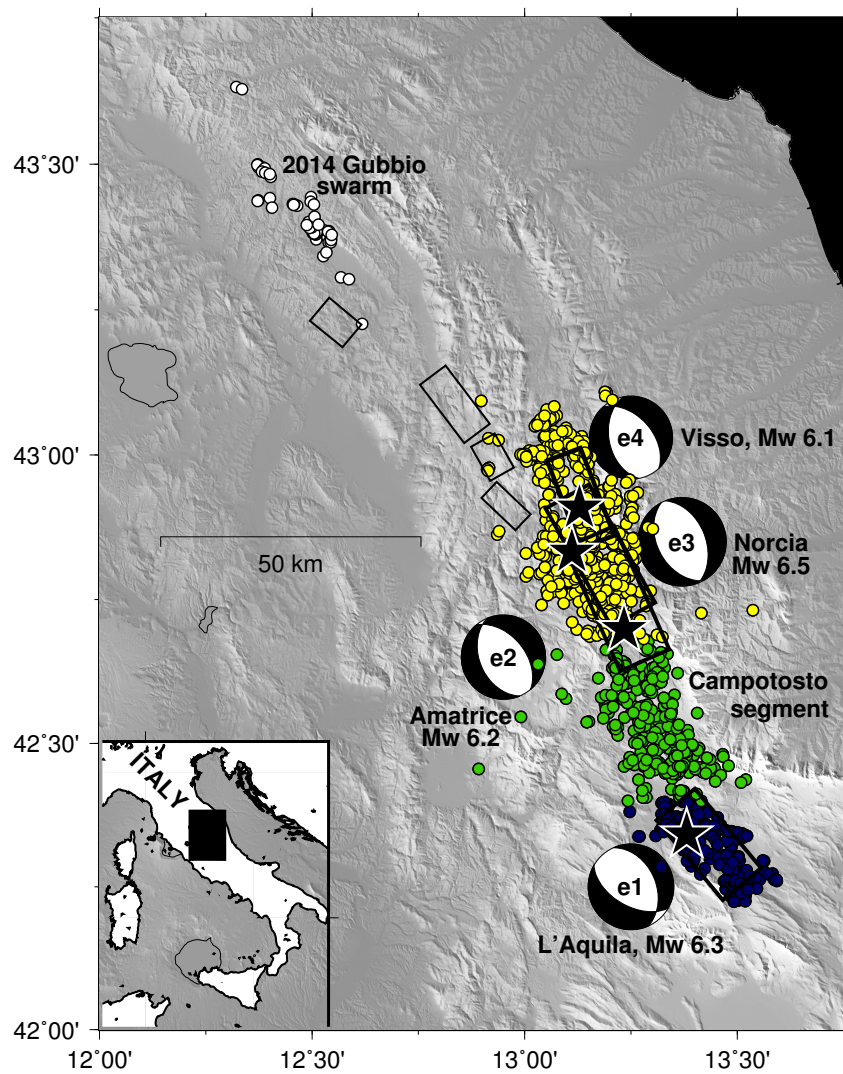


Figure 1:

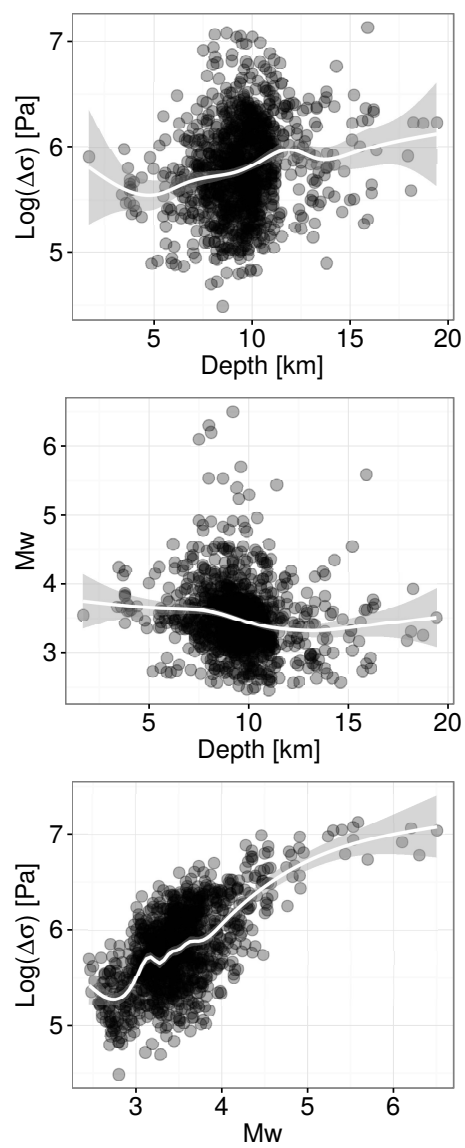


Figure 2:

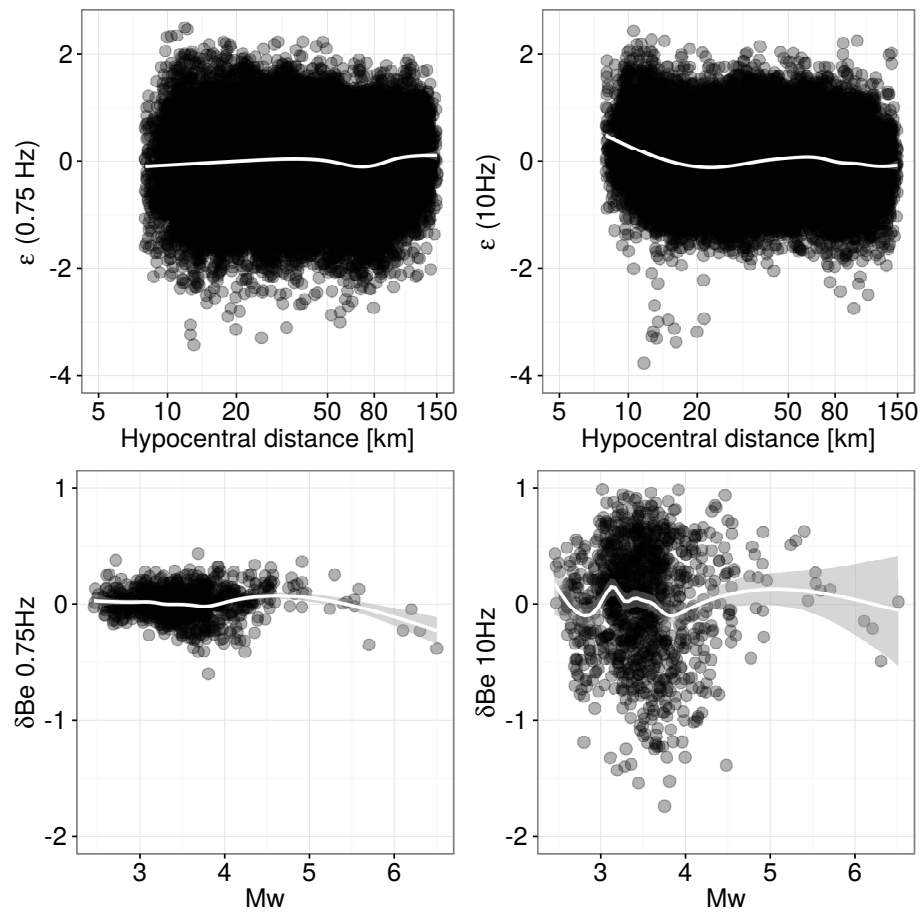


Figure 3:



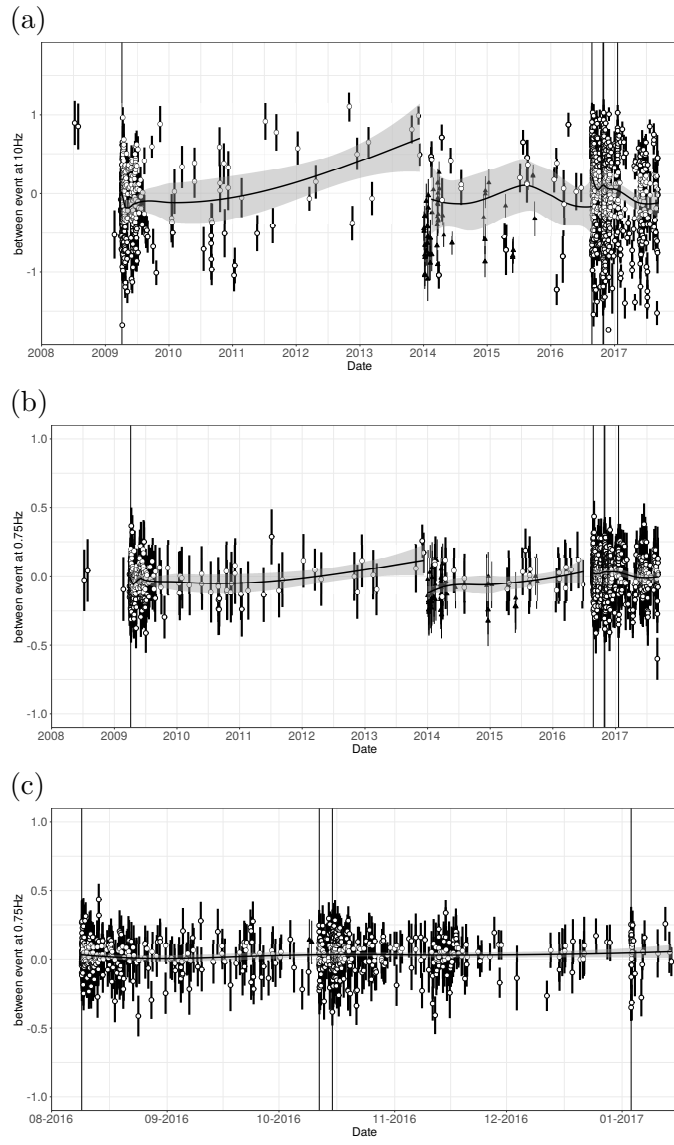


Figure 4:

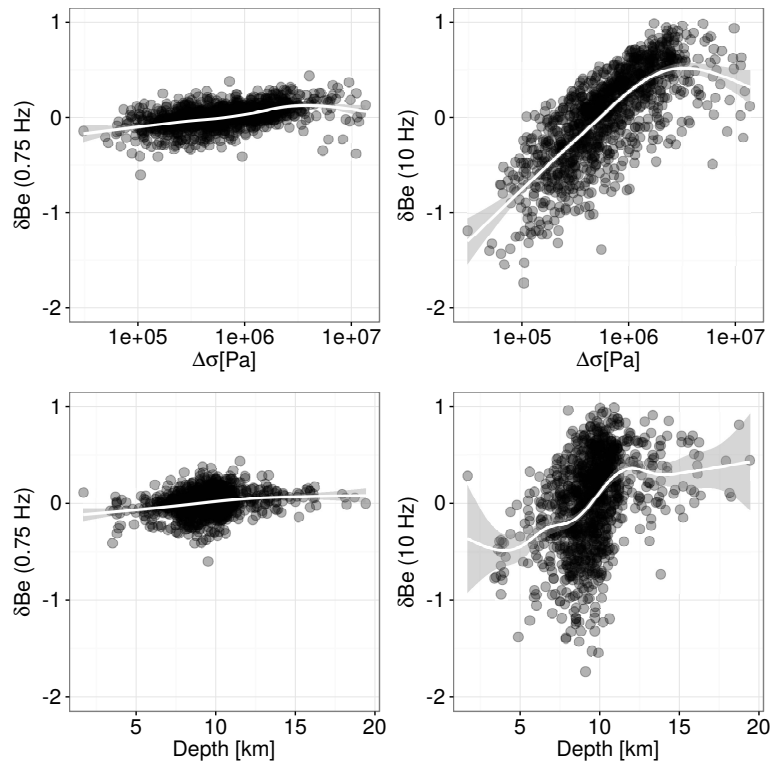


Figure 5:

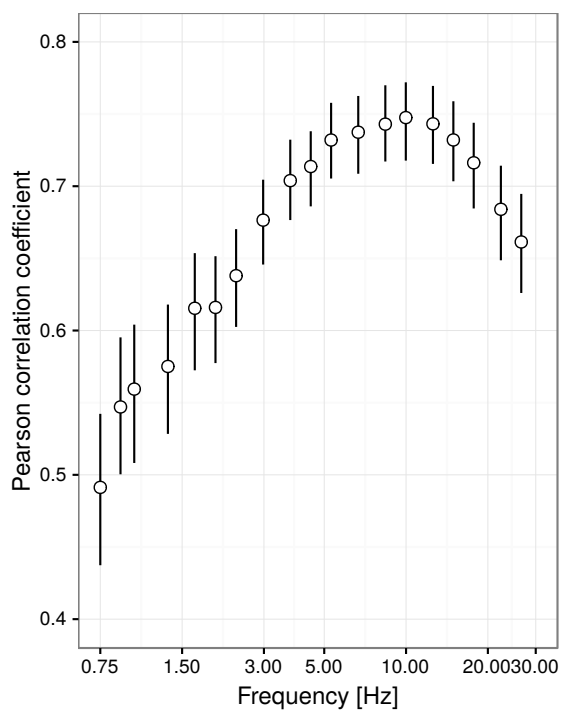


Figure 6:

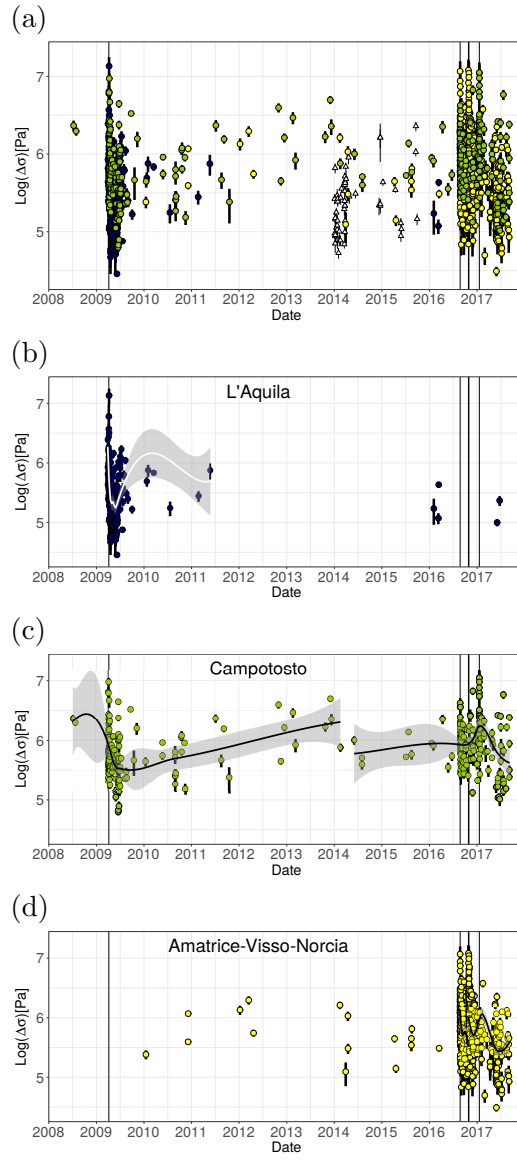


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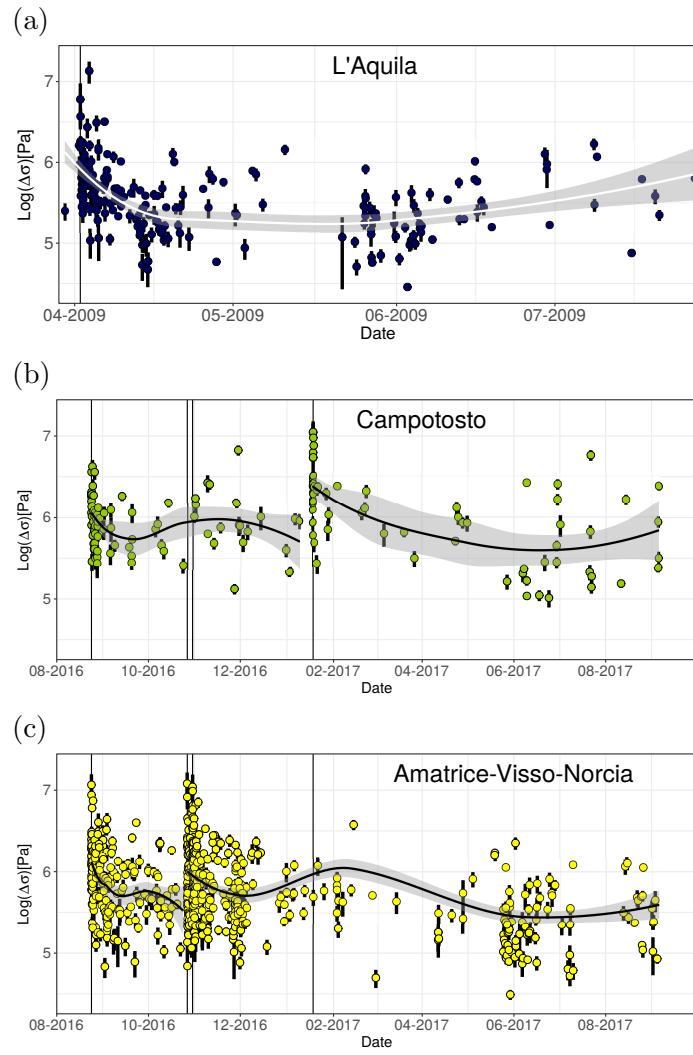
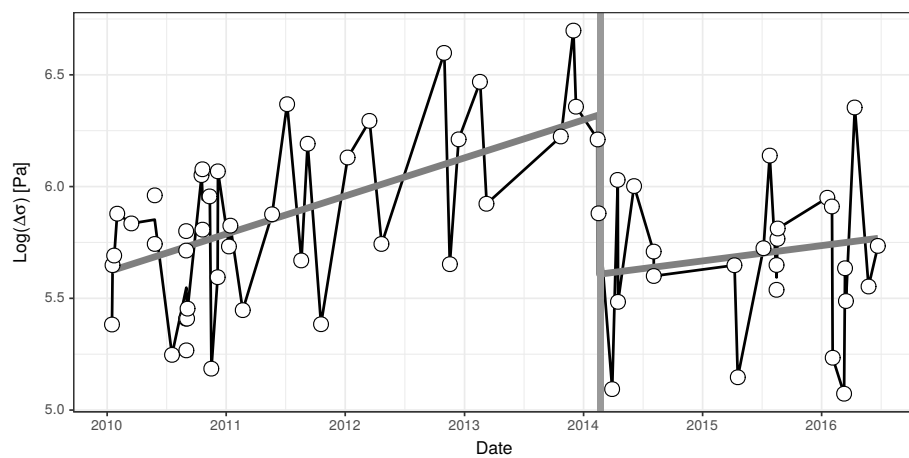


Figure 8:



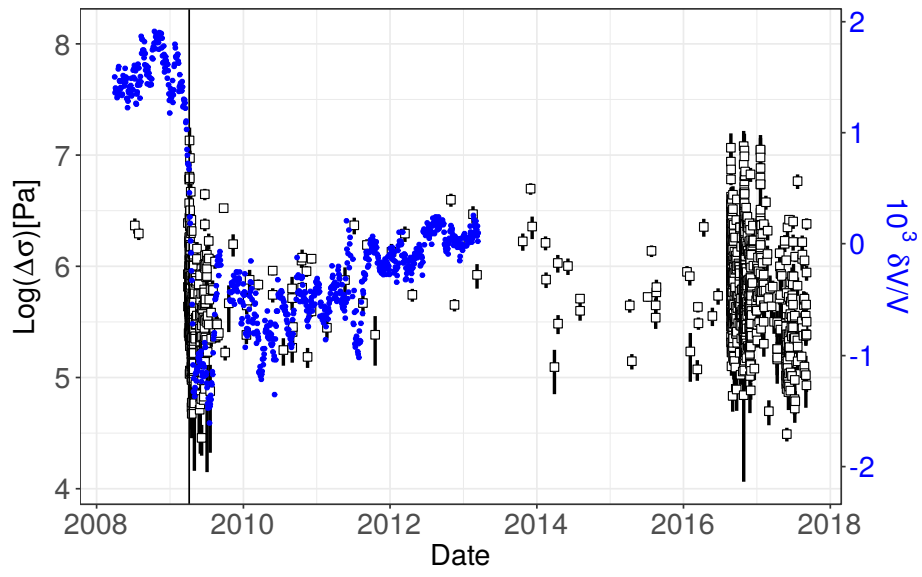


Figure 10:

## **Temporal variability of ground-shaking and stress drop in central Italy: a hint for fault healing?**

D. Bindi, F. Cotton, D. Spallarossa, M. Picozzi, and E. Rivalta

This Supplementary Material includes additional Figures cited in the main article. In figure S1, a map with the locations of stations used in this study is provided. The map in Figure S2 shows the earthquake locations as in Figure 1 but with symbols filled according to the time elapsed since the L'Aquila earthquake. Figure S3 complements Figure 4 by showing the variability with time of the between event residuals  $\delta B_e$  at 10Hz considering different time windows. In Figure S4, the symbols used to depict the time dependency of  $\delta B_e$  at 10Hz are filled according to the logarithm of the stress drop  $\Delta\sigma$ , measured in [Pa]. Finally, in Figure S5, the symbols used to show the time variability of  $\Delta\sigma$  are filled according to the moment magnitude  $M_w$ .

The Supplementary Material includes also Table S1, listing the frequency-dependent coefficients of the Ground Motion Prediction Equation developed in this study (equation 6) for the Fourier amplitude spectra. Finally, the Supplementary Material includes additional text describing both the error propagation applied for  $\Delta\sigma$  and the break-point analysis performed for locating the  $\Delta\sigma$  change-point occurred at the beginning of 2014.



## Error propagation

The non-parametric source spectra  $S(f)$  obtained through the non-parametric Generalized Inversion (Bindi et al., 2017) are described in terms of acceleration source model (Brune, 1970):

$$S_{Brune}(f) = const M_0 \frac{f^2}{1 + \left(\frac{f}{f_c}\right)^2} \quad (S1)$$

where  $M_0$  is the seismic moment,  $f_c$  the corner frequency and  $const$  is a constant depending on the radiation pattern, free surface amplification, density and velocity at the source location. We fit the  $S_{Brune}(f)$  model to  $S(f)$  by computing the logarithm (hereinafter we refer to the logarithm in base 10 as  $\log$  and to the natural logarithm as  $\ln$ ) of the spectra and considering as independent variables  $\log(M_0)$  and  $f_c$ . The fit is performed following a non-linear least-squares approach and the variance-covariance matrix  $\nu Cov$  is evaluated. From the square root of the diagonal of  $\nu Cov$ , we extracted the standard errors on  $f_c$  and on  $\log(M_0)$ , referred to as  $\delta_{f_c}$  and  $\delta_{\log M_0}$ , and from the off-diagonal elements the correlation between  $f_c$  and  $\log(M_0)$ , indicated as  $\delta_{(f_c, \log M_0)}$ . The error propagation is evaluated through the variance formula for correlated variables, truncated at the first order of the Taylor's expansion:

$$\delta_{\Delta\sigma}^2 = \left(\frac{\partial \Delta\sigma}{\partial f_c}\right)^2 \delta_{f_c}^2 + \left(\frac{\partial \Delta\sigma}{\partial \log(M_0)}\right)^2 \delta_{\log M_0}^2 + 2\delta_{(f_c, \log M_0)} \frac{\partial \Delta\sigma}{\partial f_c} \frac{\partial \Delta\sigma}{\partial \log(M_0)} \quad (S2)$$

The stress drop  $\Delta\sigma$  is computed from  $M_0$  and  $f_c$  using the following relationships (Eshelby, 1957; Brune 1970):

$$\Delta\sigma = \frac{7}{16} \frac{M_0}{r^3} \quad (S3)$$

$$r = 2.34 \frac{v_s}{2\pi f_c} \quad (S4)$$

where  $v_s$  is the shear-wave velocity at the source. Considering the logarithm of  $\Delta\sigma$ , equations (S3) and (S4) give

$$\log(\Delta\sigma) = constant + \log(M_0) + 3\log(f_c) \quad (S5)$$

and equation (2) becomes

$$\delta_{\log(\Delta\sigma)}^2 = \left(\frac{\partial \log(\Delta\sigma)}{\partial f_c}\right)^2 \delta_{f_c}^2 + \left(\frac{\partial \log(\Delta\sigma)}{\partial \log(M_0)}\right)^2 \delta_{\log M_0}^2 + 2\delta_{f_c, \log M_0} \frac{\partial \log(\Delta\sigma)}{\partial f_c} \frac{\partial \log(\Delta\sigma)}{\partial \log(M_0)} \quad (S6)$$

From equation (5),

$$\frac{\partial \log(\Delta\sigma)}{\partial \log(M_0)} = 1 \frac{\partial \log(\Delta\sigma)}{\partial f_c} = \frac{3}{\ln(10)} \frac{1}{f_c} \quad (S7)$$

Using (7) in (6), we obtain

$$\delta_{\log(\Delta\sigma)}^2 = \frac{9}{[\ln(10)]^2} \left(\frac{\delta_{f_c}}{f_c}\right)^2 + \delta_{\log M_0}^2 + \delta_{(f_c, \log M_0)} \frac{6}{\ln(10)} \frac{1}{f_c} \quad (S8)$$

The error propagation in equation (S8) is similar to the one described by Cotton et al (2013), in their equation (9). The differences are related to a different choice of the variables, since they used  $\log(f_c)$  while we used  $f_c$ , and to the fact that Cotton et al (2013) neglected the cross-term of the covariance matrix. The advantage of using  $\log(f_c)$  is that equation (S5) becomes linear with respect to the variables and, therefore, equation (S2) is exact. The drawback is that in fitting the Brune model, the dependence on the variable takes the more complex form of  $10^{\log(f_c)}$ . Regarding the cross-term, since  $\delta_{(f_c, \log M_0)}$  is negative, the choice of Cotton et al (2013) was conservative.

Finally, since  $\delta_{\Delta\sigma} = \ln(10) \Delta\sigma \delta_{\log(\Delta\sigma)}$ , from equation (S8) we get the standard error on  $\Delta\sigma$

$$\delta_{\Delta\sigma} = \Delta\sigma \sqrt{9 \left(\frac{\delta_{f_c}}{f_c}\right)^2 + [\ln(10)]^2 \delta_{\log M_0}^2 + \frac{6\ln(10)}{f_c} \delta_{(f_c, \log M_0)}} \quad (S9)$$

Figure S6 shows the 95% confidence interval, computed as 1.96 times the standard error, for the data analyzed in this study. Zooms over different time windows are shown in Figure S7.

## Breakpoint analysis

In order to assess the statistical significance of the abrupt drop in  $\Delta\sigma$  occurring at the beginning of 2014, we performed a change-point analysis (Page, 1954; Jaiswal et al., 2015). The analysis was performed through the following steps, outlined in Figure S8:

- we considered  $\Delta\sigma$  values between January 2010 and June 2016 (Figure S8a);
- since the time series is unevenly sampled, we applied an interpolation scheme. After some tests using different approaches, we used a simple linear interpolation between consecutive samples. The time series interpolated using a regular sampling of 1 day is shown in Figure S8b;
- A preliminary detection in terms of significant change in the average was performed following the approach of Killick and Eckley (2014). The detected change point (Figure S8c) is located on February 15, 2014;
- finally, the breakpoint analysis of Zeileis et al (2003) was applied to identify and locate a change in the coefficients of the linear regression with time. Figure S8d shows the location of the change point with its 95% confidence interval (corresponding to the period from February, 10 to March, 8) and the best linear models before and after the change point.

### Table caption

**Table S1.** Coefficients ( $e_1, b_1, b_2, c_1, c_2, c_3$ ) of the Fourier ground motion prediction equation described in equation (6);  $\phi_{s2s}$ ,  $\tau$ ,  $\phi_0$  are the standard deviation of the between station  $\delta Bs$ , of the between event  $\delta Be$  and residual  $\varepsilon$  distributions (see equation 6).

### Figure captions

**Figure S1** Location of stations considered in this study.

**Figure S2** Location of earthquakes considered in this study with symbols filled according to the earthquake origin time with respect to April 6, 2009 (see also Figure 1).

**Figure S3** Between-event residuals  $\delta Be$  versus time at 10 Hz for different time intervals (see also Figure 4).

**Figure S4** Between-event residuals  $\delta Be$  versus time with symbols filled according to stress drop  $\Delta\sigma$ . Triangles indicate earthquakes occurred in the area of the 2013-14 Gubbio swarm, see Figure 1.

**Figure S5** Stress drop  $\Delta\sigma$  versus time with symbols filled according to moment magnitude  $M_w$ . Triangles indicate earthquakes occurred in the area of the 2013-14 Gubbio swarm, see Figure 1.

**Figure S6.** 95% confidence intervals for  $\Delta\sigma$ .

**Figure S7.** 95% confidence intervals for  $\Delta\sigma$ , for different time intervals.

**Figure S8.** Change point analysis. a) original data. b) data linearly re-sampled at the rate of one sample per day; c) detection of the change point for the average; d) detection of the change point in the linear regression coefficients

### References

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**Table S1.** Coefficients ( $e_1, b_1, b_2, c_1, c_2, c_3$ ) of the Fourier ground motion prediction equation described in equation (6);  $\phi_{s2s}$ ,  $\tau$ ,  $\phi_0$  are the standard deviation of the between station  $\delta B_s$ , of the between event  $\delta B_e$  and residual  $\varepsilon$  distributions (see equation 6).

Freq [Hz]	$e_1$	$b_1$	$b_2$	$c_1$	$c_2$	$c_3$	$\phi_{s2s}$	$\tau$	$\phi_0$
0.35	0.80146	2.21669	-0.04089	-2.00751	0.17379	0.00545	0.47522	0.21587	0.70041
0.4	1.04779	2.27639	-0.04779	-2.07823	0.17007	0.00764	0.47742	0.20514	0.6891
0.5	1.33959	2.42340	-0.06803	-2.12578	0.15654	0.00992	0.48835	0.18558	0.64771
0.59	1.56116	2.52567	-0.1043	-2.15865	0.14717	0.01189	0.51034	0.17384	0.6542
0.75	1.92521	2.68625	-0.15401	-2.18397	0.12672	0.01265	0.54612	0.14888	0.60759
0.89	2.33067	2.71259	-0.18884	-2.23543	0.1305	0.01312	0.57532	0.14129	0.58482
1	2.60114	2.72391	-0.21177	-2.27032	0.12893	0.01357	0.59266	0.14381	0.59366
1.33	3.09564	2.76270	-0.23556	-2.26438	0.10171	0.01225	0.61237	0.16665	0.56748
1.67	3.48468	2.68448	-0.24237	-2.27198	0.0962	0.01163	0.61487	0.19459	0.53169
1.99	3.67784	2.62330	-0.24584	-2.26209	0.0945	0.01103	0.6234	0.22036	0.53082
2.37	3.76324	2.55504	-0.23485	-2.20498	0.08715	0.00903	0.63017	0.25424	0.52461
2.98	3.59658	2.38582	-0.21798	-2.04151	0.0927	0.00468	0.60396	0.29829	0.52005
3.75	3.42988	2.25625	-0.19759	-1.90391	0.08802	0.00061	0.62728	0.34572	0.52529
4.46	3.28532	2.08639	-0.17907	-1.79832	0.10474	-0.00292	0.65699	0.37856	0.5284
5.3	3.07226	2.00030	-0.15746	-1.69606	0.09872	-0.00611	0.69546	0.40803	0.52817
6.67	2.87406	1.86515	-0.13773	-1.60836	0.10035	-0.00995	0.76182	0.44405	0.52234
8.39	2.57230	1.78698	-0.11523	-1.51731	0.08999	-0.01435	0.81647	0.48376	0.52027
10	2.36854	1.71378	-0.09805	-1.47415	0.08535	-0.01735	0.87734	0.51209	0.51604
12.56	2.31993	1.61380	-0.08088	-1.50875	0.08424	-0.02057	0.96463	0.56033	0.51112
14.93	2.34572	1.50170	-0.06593	-1.57316	0.09194	-0.02223	102.517	0.59434	0.51074
17.74	2.35082	1.44668	-0.03929	-1.67057	0.08297	-0.0224	108.441	0.63308	0.5131
22.33	2.52609	1.39944	-0.01618	-1.91863	0.06611	-0.01832	112.420	0.68558	0.52859
26.54	2.67305	1.39498	0.01618	-2.11285	0.04734	-0.01503	109.470	0.70831	0.52971

